

SPECIAL DISCRETE RANDOM VARIABLES

The Bernoulli Random Variable

A *Bernoulli trial* has only *two outcomes*, with probability

$$P(X = 1) = p ,$$

$$P(X = 0) = 1 - p ,$$

e.g., tossing a coin, winning or losing a game, \dots .

We have

$$E[X] = 1 \cdot p + 0 \cdot (1 - p) = p ,$$

$$E[X^2] = 1^2 \cdot p + 0^2 \cdot (1 - p) = p ,$$

$$\text{Var}(X) = E[X^2] - E[X]^2 = p - p^2 = p(1 - p) .$$

NOTE : If p is small then $\text{Var}(X) \cong p$.

EXAMPLES :

- When $p = \frac{1}{2}$ (e.g., for tossing a coin), we have

$$E[X] = p = \frac{1}{2} \quad , \quad Var(X) = p(1 - p) = \frac{1}{4} .$$

- When *rolling a die* , with outcome k , ($1 \leq k \leq 6$) , let

$$X(k) = 1 \quad \text{if the roll resulted in a } \textit{six} ,$$

and

$$X(k) = 0 \quad \text{if the roll did } \textit{not} \text{ result in a } \textit{six} .$$

Then

$$E[X] = p = \frac{1}{6} \quad , \quad Var(X) = p(1 - p) = \frac{5}{36} .$$

- When $p = 0.01$, then

$$E[X] = 0.01 \quad , \quad Var(X) = 0.0099 \cong 0.01 .$$

The Binomial Random Variable

Perform a Bernoulli trial n times *in sequence* .

Assume the individual trials are *independent* .

An *outcome* could be

$$100011001010 \quad (n = 12) ,$$

with probability

$$P(100011001010) = p^5 \cdot (1 - p)^7 . \quad (\text{Why ?})$$

Let the X be the number of "*successes*" (*i.e.* 1's) .

For example,

$$X(100011001010) = 5 .$$

We have

$$P(X = 5) = \binom{12}{5} \cdot p^5 \cdot (1 - p)^7 . \quad (\text{Why ?})$$

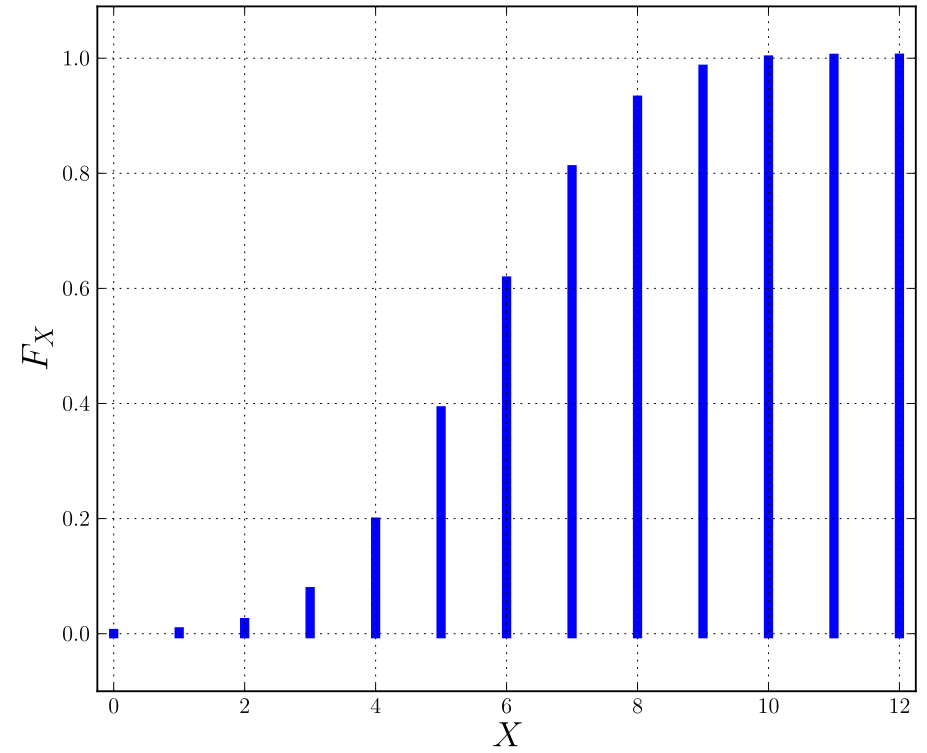
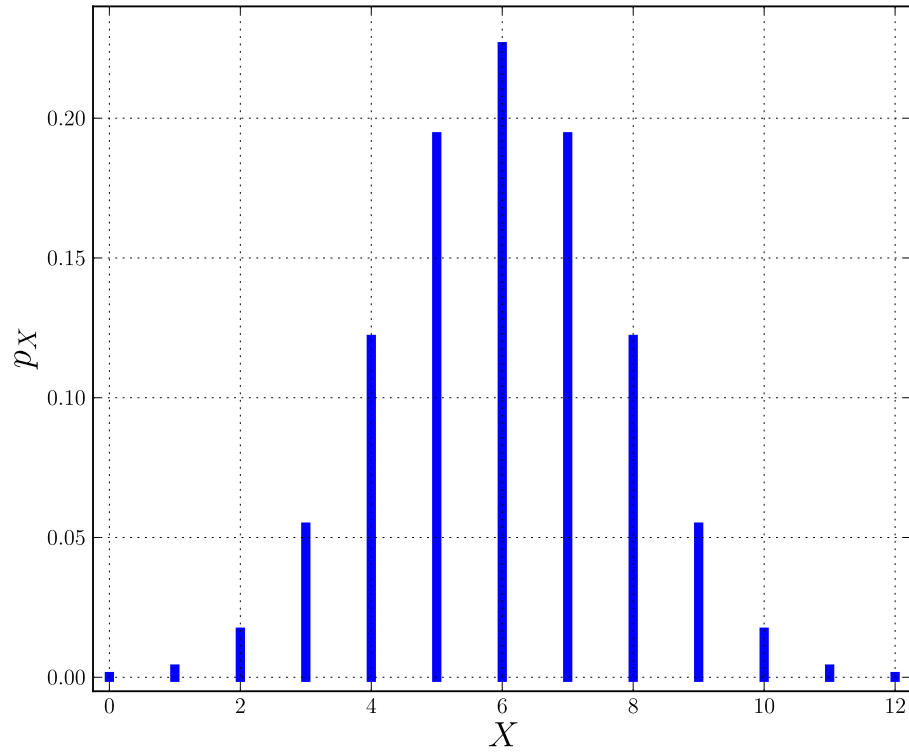
In general, for k successes in a sequence of n trials, we have

$$P(X = k) = \binom{n}{k} \cdot p^k \cdot (1 - p)^{n-k}, \quad (0 \leq k \leq n).$$

EXAMPLE : Tossing a coin 12 times:

$$n = 12, \quad p = \frac{1}{2}$$

k	$p_X(k)$	$F_X(k)$
0	1 / 4096	1 / 4096
1	12 / 4096	13 / 4096
2	66 / 4096	79 / 4096
3	220 / 4096	299 / 4096
4	495 / 4096	794 / 4096
5	792 / 4096	1586 / 4096
6	924 / 4096	2510 / 4096
7	792 / 4096	3302 / 4096
8	495 / 4096	3797 / 4096
9	220 / 4096	4017 / 4096
10	66 / 4096	4083 / 4096
11	12 / 4096	4095 / 4096
12	1 / 4096	4096 / 4096



The Binomial *mass* and *distribution* functions for $n = 12$, $p = \frac{1}{2}$

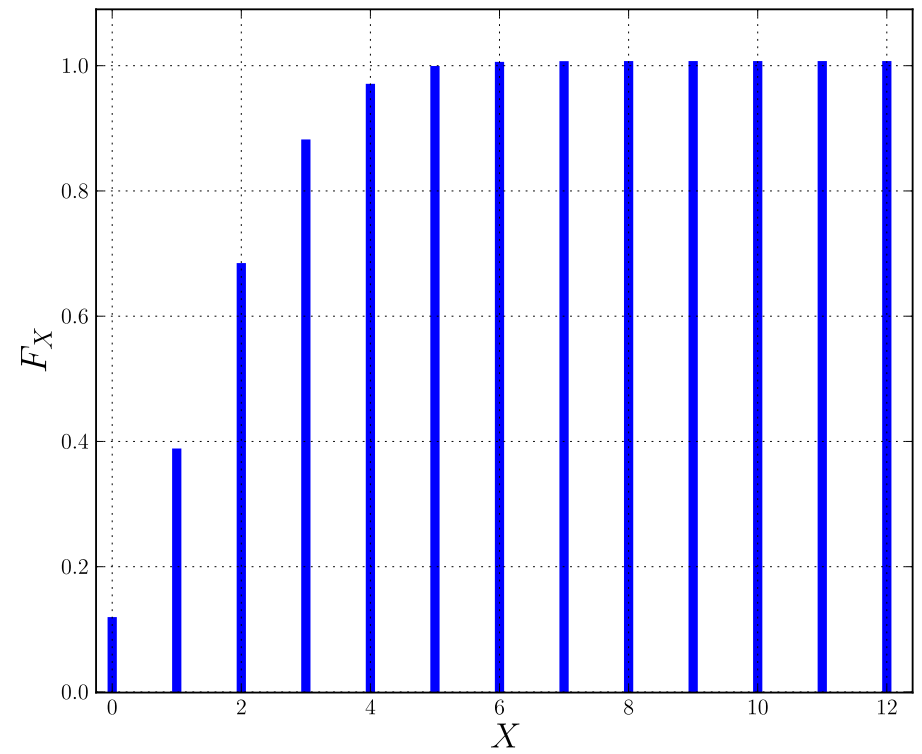
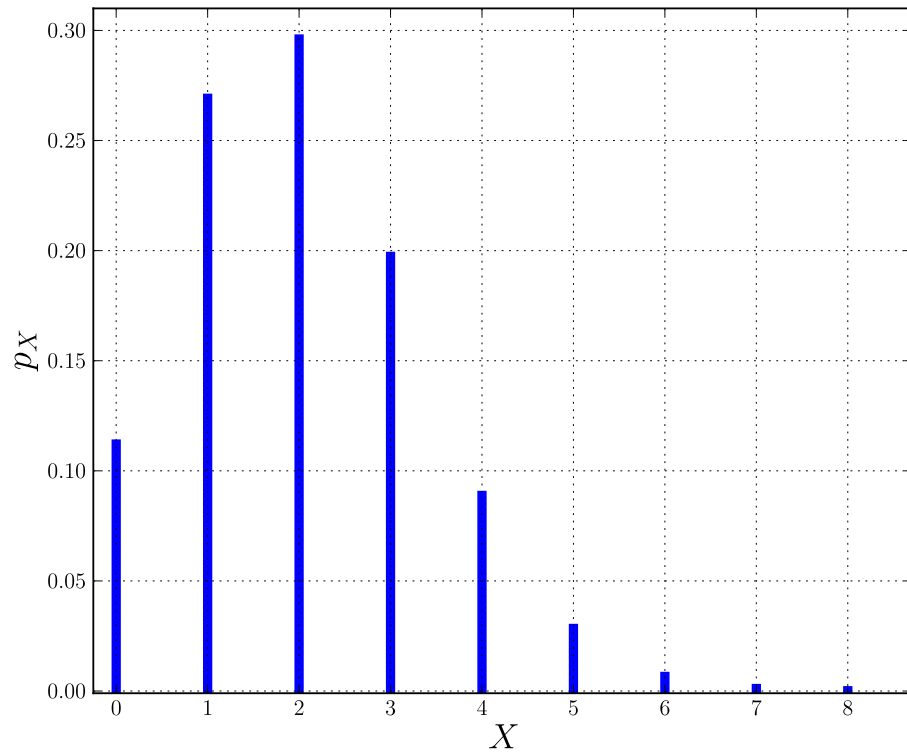
For k successes in a sequence of n trials :

$$P(X = k) = \binom{n}{k} \cdot p^k \cdot (1 - p)^{n-k}, \quad (0 \leq k \leq n).$$

EXAMPLE : Rolling a die 12 times:

$$n = 12, \quad p = \frac{1}{6}$$

k	$p_X(k)$	$F_X(k)$
0	0.1121566221	0.112156
1	0.2691758871	0.381332
2	0.2960935235	0.677426
3	0.1973956972	0.874821
4	0.0888280571	0.963649
5	0.0284249838	0.992074
6	0.0066324966	0.998707
7	0.0011369995	0.999844
8	0.0001421249	0.999986
9	0.0000126333	0.999998
10	0.0000007580	0.999999
11	0.0000000276	0.999999
12	0.0000000005	1.000000



The Binomial *mass* and *distribution* functions for $n = 12$, $p = \frac{1}{6}$

EXAMPLE :

In 12 *rolls of a die* write the outcome as, for example,

100011001010

where

1 denotes the roll resulted in a *six* ,

and

0 denotes the roll did *not* result in a *six* .

As before, let X be the number of 1's in the outcome.

Then X represents the *number of sixes* in the 12 rolls.

Then, for example, using the preceding *Table* :

$$P(X = 5) \cong 2.8 \% \quad , \quad P(X \leq 5) \cong 99.2 \% .$$

EXERCISE : Show that from

$$P(X = k) = \binom{n}{k} \cdot p^k \cdot (1 - p)^{n-k} ,$$

and

$$P(X = k + 1) = \binom{n}{k + 1} \cdot p^{k+1} \cdot (1 - p)^{n-k-1} ,$$

it follows that

$$P(X = k + 1) = c_k \cdot P(X = k) ,$$

where

$$c_k = \frac{n - k}{k + 1} \cdot \frac{p}{1 - p} .$$

NOTE : This *recurrence formula* is an efficient and stable *algorithm* to compute the binomial probabilities :

$$P(X = 0) = (1 - p)^n ,$$

$$P(X = k + 1) = c_k \cdot P(X = k) , \quad k = 0, 1, \dots, n - 1 .$$

Mean and variance of the Binomial random variable :

By definition, the *mean* of a Binomial random variable X is

$$E[X] = \sum_{k=0}^n k \cdot P(X = k) = \sum_{k=0}^n k \cdot \binom{n}{k} p^k (1-p)^{n-k},$$

which can be shown to equal np .

An *easy way* to see this is as follows :

If in a *sequence* of n independent Bernoulli trials we let

$X_k =$ the outcome of the k^{th} Bernoulli trial , $(X_k = 0 \text{ or } 1)$,

then

$$X \equiv X_1 + X_2 + \cdots + X_n ,$$

is the *Binomial random variable* that *counts the successes* ” .

$$X \equiv X_1 + X_2 + \cdots + X_n$$

We know that

$$E[X_k] = p ,$$

so

$$E[X] = E[X_1] + E[X_2] + \cdots + E[X_n] = np .$$

We already know that

$$\text{Var}(X_k) = E[X_k^2] - (E[X_k])^2 = p - p^2 = p(1 - p) ,$$

so, since the X_k are *independent*, we have

$$\text{Var}(X) = \text{Var}(X_1) + \text{Var}(X_2) + \cdots + \text{Var}(X_n) = np(1 - p) .$$

NOTE : If p is small then $\text{Var}(X) \cong np$.

EXAMPLES :

- For 12 tosses of a *coin* , with *Heads* is *success*, we have

so
$$n = 12 \quad , \quad p = \frac{1}{2}$$

$$E[X] = np = 6 \quad , \quad \text{Var}(X) = np(1 - p) = 3 .$$

- For 12 rolls of a *die* , with *six* is *success* , we have

so
$$n = 12 \quad , \quad p = \frac{1}{6}$$

$$E[X] = np = 2 \quad , \quad \text{Var}(X) = np(1 - p) = 5/3 .$$

- If $n = 500$ and $p = 0.01$, then

$$E[X] = np = 5 \quad , \quad \text{Var}(X) = np(1 - p) = 4.95 \cong 5 .$$

The Poisson Random Variable

The Poisson variable *approximates* the Binomial random variable :

$$P(X = k) = \binom{n}{k} \cdot p^k \cdot (1 - p)^{n-k} \cong e^{-\lambda} \cdot \frac{\lambda^k}{k!} ,$$

when we take

$$\lambda = n p \quad (\text{the average number of successes}) .$$

This approximation is *accurate* if n is *large* and p *small* .

Recall that for the **Binomial** random variable

$$E[X] = n p , \text{ and } Var(X) = np(1 - p) \cong np \text{ when } p \text{ is small.}$$

Indeed, for the **Poisson** random variable we will show that

$$E[X] = \lambda \quad \text{and} \quad Var(X) = \lambda .$$

A *stable* and *efficient* way to compute the Poisson probability

$$P(X = k) = e^{-\lambda} \cdot \frac{\lambda^k}{k!}, \quad k = 0, 1, 2, \dots,$$

$$P(X = k + 1) = e^{-\lambda} \cdot \frac{\lambda^{k+1}}{(k + 1)!},$$

is to use the *recurrence relation*

$$P(X = 0) = e^{-\lambda},$$

$$P(X = k + 1) = \frac{\lambda}{k + 1} \cdot P(X = k), \quad k = 0, 1, 2, \dots.$$

NOTE : Unlike the Binomial random variable, the Poisson random variable can have an *arbitrarily large* integer value k .

The Poisson random variable

$$P(X = k) = e^{-\lambda} \cdot \frac{\lambda^k}{k!}, \quad k = 0, 1, 2, \dots,$$

has (as shown later) : $E[X] = \lambda$ and $Var(X) = \lambda$.

The Poisson *distribution function* is

$$F(k) = P(X \leq k) = \sum_{\ell=0}^k e^{-\lambda} \frac{\lambda^\ell}{\ell!} = e^{-\lambda} \sum_{\ell=0}^k \frac{\lambda^\ell}{\ell!},$$

with, as should be the case,

$$\lim_{k \rightarrow \infty} F(k) = e^{-\lambda} \sum_{\ell=0}^{\infty} \frac{\lambda^\ell}{\ell!} = e^{-\lambda} e^{\lambda} = 1.$$

(using the *Taylor series* from Calculus for e^{λ}).

The Poisson random variable

$$P(X = k) = e^{-\lambda} \cdot \frac{\lambda^k}{k!}, \quad k = 0, 1, 2, \dots,$$

models the probability of k "successes" in a given "time" interval, when the *average* number of successes is λ .

EXAMPLE : Suppose customers arrive at the rate of *six* per hour. The probability that k customers arrive in a one-hour period is

$$P(k = 0) = e^{-6} \cdot \frac{6^0}{0!} \cong 0.0024,$$

$$P(k = 1) = e^{-6} \cdot \frac{6^1}{1!} \cong 0.0148,$$

$$P(k = 2) = e^{-6} \cdot \frac{6^2}{2!} \cong 0.0446.$$

The probability that more than 2 customers arrive is

$$1 - (0.0024 + 0.0148 + 0.0446) \cong 0.938.$$

$$p_{\text{Binomial}}(k) = \binom{n}{k} p^k (1-p)^{n-k} \cong p_{\text{Poisson}}(k) = e^{-\lambda} \frac{\lambda^k}{k!}$$

EXAMPLE : $\lambda = 6$ customers/hour.

For the Binomial take $n = 12$, $p = 0.5$ (0.5 customers/5 minutes),
so that indeed $np = \lambda$.

k	p_{Binomial}	p_{Poisson}	F_{Binomial}	F_{Poisson}
0	0.0002	0.0024	0.0002	0.0024
1	0.0029	0.0148	0.0031	0.0173
2	0.0161	0.0446	0.0192	0.0619
3	0.0537	0.0892	0.0729	0.1512
4	0.1208	0.1338	0.1938	0.2850
5	0.1933	0.1606	0.3872	0.4456
6	0.2255	0.1606	0.6127	0.6063
7	0.1933	0.1376	0.8061	0.7439
8	0.1208	0.1032	0.9270	0.8472
9	0.0537	0.0688	0.9807	0.9160
10	0.0161	0.0413	0.9968	0.9573
11	0.0029	0.0225	0.9997	0.9799
12	0.0002	0.0112	1.0000	0.9911★

Why not 1.0000 ?

Here the approximation is *not so good* ...

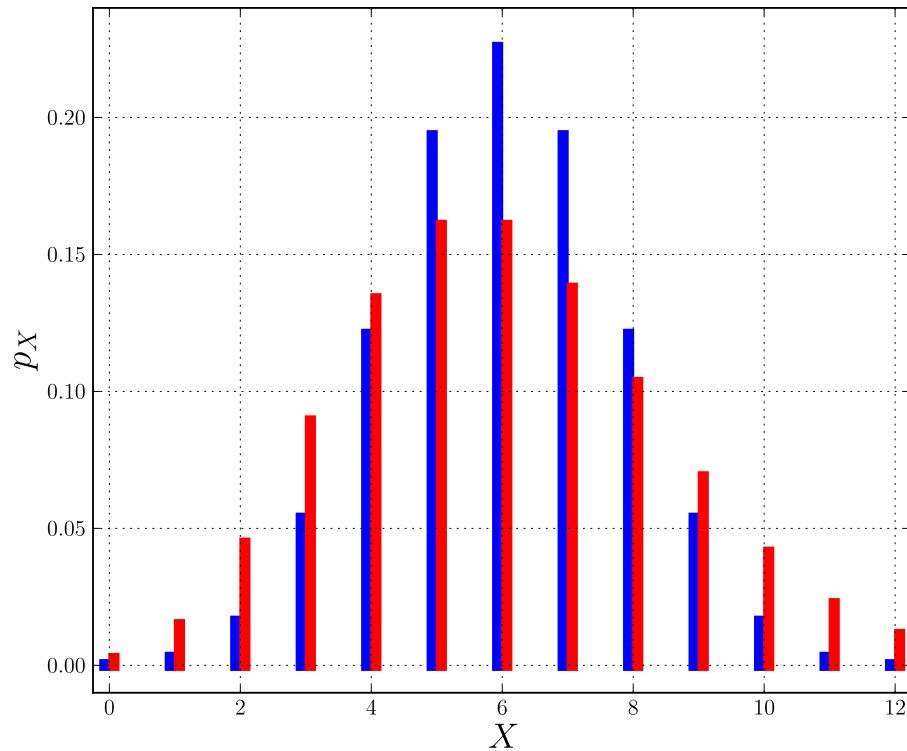
$$p_{\text{Binomial}}(k) = \binom{n}{k} p^k (1-p)^{n-k} \cong p_{\text{Poisson}}(k) = e^{-\lambda} \frac{\lambda^k}{k!}$$

EXAMPLE : $\lambda = 6$ customers/hour.

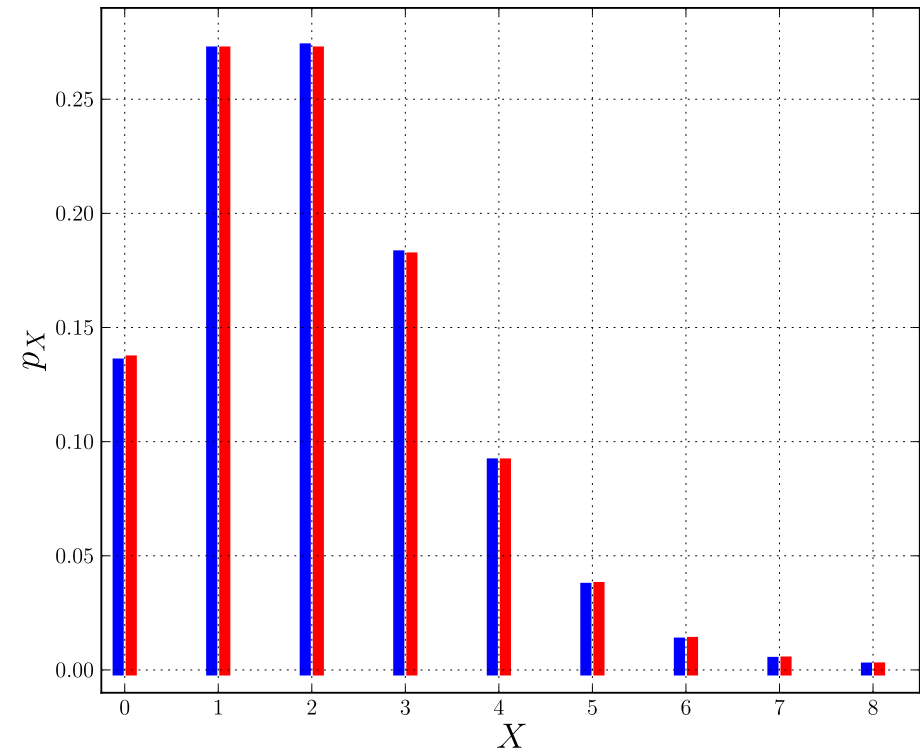
For the Binomial take $n = 60$, $p = 0.1$ (0.1 customers/minute) ,
so that indeed $np = \lambda$.

k	p_{Binomial}	p_{Poisson}	F_{Binomial}	F_{Poisson}
0	0.0017	0.0024	0.0017	0.0024
1	0.0119	0.0148	0.0137	0.0173
2	0.0392	0.0446	0.0530	0.0619
3	0.0843	0.0892	0.1373	0.1512
4	0.1335	0.1338	0.2709	0.2850
5	0.1662	0.1606	0.4371	0.4456
6	0.1692	0.1606	0.6064	0.6063
7	0.1451	0.1376	0.7515	0.7439
8	0.1068	0.1032	0.8583	0.8472
9	0.0685	0.0688	0.9269	0.9160
10	0.0388	0.0413	0.9657	0.9573
11	0.0196	0.0225	0.9854	0.9799
12	0.0089	0.0112	0.9943	0.9911
13

Here the approximation is *better* ...



$$n = 12 \quad , \quad p = \frac{1}{2} \quad , \quad \lambda = 6$$



$$n = 200 \quad , \quad p = 0.01 \quad , \quad \lambda = 2$$

The Binomial (*blue*) and Poisson (*red*) probability mass functions.

For the case $n = 200$, $p = 0.01$, the approximation is very good !

For the *Binomial* random variable we found

$$E[X] = np \quad \text{and} \quad \text{Var}(X) = np(1 - p) ,$$

while for the *Poisson* random variable, with $\lambda = np$ we will show

$$E[X] = np \quad \text{and} \quad \text{Var}(X) = np .$$

Note again that

$$np(1 - p) \cong np , \quad \text{when } p \text{ is } \textit{small} .$$

EXAMPLE : In the preceding two *Tables* we have

$$n=12 , \quad p=0.5$$

	Binomial	Poisson
$E[X]$	6.0000	6.0000
$\text{Var}[X]$	3.0000	6.0000
$\sigma[X]$	1.7321	2.4495

$$n=60 , \quad p=0.1$$

	Binomial	Poisson
$E[X]$	6.0000	6.0000
$\text{Var}[X]$	5.4000	6.0000
$\sigma[X]$	2.3238	2.4495

FACT : (*The Method of Moments*)

By *Taylor expansion* of e^{tX} about $t = 0$, we have

$$\begin{aligned}\psi(t) &\equiv E[e^{tX}] = E\left[1 + tX + \frac{t^2 X^2}{2!} + \frac{t^3 X^3}{3!} + \dots\right] \\ &= 1 + t E[X] + \frac{t^2}{2!} E[X^2] + \frac{t^3}{3!} E[X^3] + \dots.\end{aligned}$$

It follows that

$$\psi'(0) = E[X] \quad , \quad \psi''(0) = E[X^2] . \quad (\text{Why ?})$$

This sometimes *facilitates computing the mean*

$$\mu = E[X] ,$$

and the variance

$$\text{Var}(X) = E[X^2] - \mu^2 .$$

APPLICATION : The *Poisson mean* and *variance* :

$$\begin{aligned}\psi(t) &\equiv E[e^{tX}] = \sum_{k=0}^{\infty} e^{tk} P(X = k) = \sum_{k=0}^{\infty} e^{tk} e^{-\lambda} \frac{\lambda^k}{k!} \\ &= e^{-\lambda} \sum_{k=0}^{\infty} \frac{(\lambda e^t)^k}{k!} = e^{-\lambda} e^{\lambda e^t} = e^{\lambda(e^t - 1)} .\end{aligned}$$

Here $\psi'(t) = \lambda e^t e^{\lambda(e^t - 1)}$

$$\psi''(t) = \lambda [\lambda (e^t)^2 + e^t] e^{\lambda(e^t - 1)} \quad (\text{Check !})$$

so that

$$E[X] = \psi'(0) = \lambda$$

$$E[X^2] = \psi''(0) = \lambda(\lambda + 1) = \lambda^2 + \lambda$$

$$\text{Var}(X) = E[X^2] - E[X]^2 = \lambda .$$

EXAMPLE : *Defects* in a wire occur at the rate of *one per 10 meter*, with a *Poisson distribution* :

$$P(X = k) = e^{-\lambda} \cdot \frac{\lambda^k}{k!}, \quad k = 0, 1, 2, \dots .$$

What is the probability that :

- A 12-meter roll has at *no* defects?

ANSWER : Here $\lambda = 1.2$, and $P(X = 0) = e^{-\lambda} = 0.3012$.

- A 12-meter roll of wire has *one* defect?

ANSWER : With $\lambda = 1.2$, $P(X = 1) = e^{-\lambda} \cdot \lambda = 0.3614$.

- Of *five* 12-meter rolls *two* have *one* defect and *three* have *none*?

ANSWER : $\binom{5}{3} \cdot 0.3012^3 \cdot 0.3614^2 = 0.0357$. (**Why ?**)

EXERCISE :

Defects in a certain wire occur at the rate of one per 10 meter.

Assume the defects have a Poisson distribution.

What is the probability that :

- a 20-meter wire has no defects?
- a 20-meter wire has at most 2 defects?

EXERCISE :

Customers arrive at a counter at the rate of 8 per hour.

Assume the arrivals have a Poisson distribution.

What is the probability that :

- no customer arrives in 15 minutes?
- two customers arrive in a period of 30 minutes?